Three challenges for spatiotemporal Hawkes modeling

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September 28, 2021

Spatiotemporal data in public health

Washington D.C. gunshots (2018)





2,925 Wildfire ignition sites in Alaska: 2015–2019

Global influenza (2000-2012)



Poisson processes

A counting process $\{N(t), t > 0\}$ is a homogeneous Poisson point process with rate $\lambda > 0$ if

(i)
$$N(0) = 0$$
;
(ii) $(N(t_4) - N(t_3)) \perp (N(t_2) - N(t_1))$ for $t_1 < t_2 \le t_3 < t_4$;
(iii) $(N(t_2) - N(t_1)) \sim Poisson(\lambda(t_2 - t_1))$ for $t_2 > t_1$.

It is an inhomogeneous Poisson point process with rate $\lambda(t) > 0$ if (i) N(0) = 0; (ii) $(N(t_4) - N(t_3)) \perp (N(t_2) - N(t_1))$ for $t_1 < t_2 < t_3 < t_4$;

(iii)
$$(N(t_2) - N(t_1)) \sim Poisson(\int_{t_1}^{t_2} \lambda(t) dt)$$
 for $t_2 > t_1$.

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Hawkes process



$$\lambda(t) = \mu + \xi(t) = \mu + \sum_{t_n < t} g(t - t_n)$$

10

Spatiotemporal Hawkes process



$$\lambda(\mathsf{x},t) = \mu(\mathsf{x}) + \xi(\mathsf{x},t) = \mu(\mathsf{x}) + \sum_{t_n < t} g(\mathsf{x} - \mathsf{x}_n, t - t_n)$$

A simple model

We assume

- 1. an exponential decay triggering function,
- 2. Gaussian kernel spatial smoothers, and
- 3. separability in space/time:

$$\xi(\mathbf{x},t) = \frac{\theta_0 \omega}{h^D} \sum_{t_n < t} e^{-\omega (t-t_n)} \phi\left(\frac{\mathbf{x} - \mathbf{x}_n}{h}\right)$$
$$\mu(\mathbf{x},t) = \frac{\mu_0}{\tau_x^D \tau_t} \sum_{n=1}^N \phi\left(\frac{\mathbf{x} - \mathbf{x}_n}{\tau_x}\right) \cdot \phi\left(\frac{t-t_n}{\tau_t}\right) \,.$$

Three inferential challenges

Likelihood based inference encounters (at least) three challenges that are not independent from one another.

big data \times spatial data precision \times big model

Big data

Washington D.C. gunshots (2018)



D.C. gunshot data (2006-2018)

An acoustic gunshot location system recorded over 85k gunshots in Washington D.C. between 2006 and 2018.

Loeffler and Flaxman (2018) used a subset of 9k gunshots in the paper titled *Is gun violence contagious? A spatiotemporal test.*

They answered 'yes', but did the results hold for a complete data analysis?

Likelihood based inference

The likelihood for data (x_1, t_1) , ..., (x_N, t_N) is

$$\mathcal{L}(\Theta) = \exp\left(-\int_{\mathbb{R}^D} \int_0^{t_N} \lambda(\mathbf{x}, t) \, \mathrm{d}t \, \mathrm{d}\mathbf{x}\right) \prod_{n=1}^N \lambda(\mathbf{x}_n, t_n)$$
$$:= e^{-\Lambda(t_N)} \cdot \prod_{n=1}^N \lambda_n.$$

The log-likelihood involves the term

$$\sum_{n=1}^{N} \log \lambda_n = \sum_{n=1}^{N} \log \left(\mu_n + \frac{\theta_0 \omega}{h^D} \sum_{t_{n'} < t_n} e^{-\omega (t_n - t_{n'})} \phi \left(\frac{\mathsf{x}_n - \mathsf{x}_{n'}}{h} \right) \right)$$

The gradient w.r.t. Θ also features a double summation.

Parallelization methods

Central processing unit (CPU):

- 1. Global parallelization: 2 to hundreds of cores (multi-core)
- 2. Local parallelization: single instruction multiple data (SIMD)

Graphics processing unit (GPU):

- 1. Thousands of cores (many-core)
- 2. Single instruction multiple threads (SIMT)
- 3. High memory bandwidth

Significant speedups



Significant speedups

Seconds per evaluation			Relative	e speedu	р
GPU	C++	R	GPU	C++	R
0.004	0.80	5.02	1255.00	6.27	1
0.01	2.66	18.74	1338.57	7.05	1
0.05	10.10	105.54	1991.32	10.45	1
0.12	21.10	232.51	1970.42	11.02	1
	Second GPU 0.004 0.01 0.05 0.12	Seconds per ev GPU C+++ 0.004 0.80 0.01 2.66 0.05 10.10 0.12 21.10	Seconds per evaluation GPU C++ R 0.004 0.80 5.02 0.01 2.66 18.74 0.05 10.10 105.54 0.12 21.10 232.51	Seconds per evaluation Relative GPU C++ R GPU 0.004 0.80 5.02 1255.00 0.01 2.66 18.74 1338.57 0.05 10.10 105.54 1991.32 0.12 21.10 232.51 1970.42	Seconds per evaluation Relative speedu GPU C++ R GPU C++ 0.004 0.80 5.02 1255.00 6.27 0.01 2.66 18.74 1338.57 7.05 0.05 10.10 105.54 1991.32 10.45 0.12 21.10 232.51 1970.42 11.02



Postprocessing is expensive too

We can also consider the posterior distribution for the probability an event comes from self-excitation: $\xi_n/(\xi_n + \mu_n)$.



Spatial data precision

Washington D.C. gunshots (2018)





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Global influenza (2000-2012)



Simultaneously inferring gunshot locations

The P.D. rounds the data to the nearest 100 meters. A uniform prior over the 10k m^2 square centered at each observation \mathfrak{x}_n

$$p(\mathsf{x}_n) \propto 1$$
, $\mathfrak{x}_{nd} - 50 < \mathsf{x}_{nd} < \mathfrak{x}_{nd} + 50$, $d = 1, 2$

corresponds to using the *grouped data likelihood* of Heitjan and Rubin (1991).

Simultaneously inferring gunshot locations



Simultaneously inferring gunshot locations



Good news!



Posterior median (95%)	Credible interval)
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Rate component	Parameter	Full model	Naive model
Background	Spatial lengthscale (m)	98.1 (94.0, 103.3)	106.3 (102.1, 110.7)
	Temporal lengthscale (hrs)	1763.7 (1552.9, 2014.8)	1891.8 (1665.1, 2163.6)
Self-excitatory	Spatial lengthscale (m)	61.4 (56.4, 67.2)	72.3 (67.9, 77.2)
	Temporal lengthscale (hrs)	0.009 (0.008, 0.010)	0.009 (0.008, 0.009)
	Normalized weight	0.11 (0.10, 0.12)	0.11 (0.10, 0.12)

Breaking the model: decreasing precision



	50% Cls			80% Cls		95% Cls			
Spatial precision	1.0	0.5	0.1	1.0	0.5	0.1	1.0	0.5	0.1
Fixed locations	0.00	0.19	0.52	0.00	0.42	0.81	0.00	0.68	0.96
Sampled locations	0.53	0.49	0.53	0.84	0.81	0.81	0.98	0.95	0.96

Breaking the model: variable precision

2,925 Wildfire ignition sites in Alaska: 2015-2019



Breaking the model: variable precision



Rate component	Parameter	Full model	Naive model A	Naive model B
Background	Spatial lengthscale (km)	34.8 (32.9, 37.6)	23.5 (22.3, 24.6)	63.0 (58.7, 68.7)
	Temporal lengthscale (days)	25.9 (23.8, 27.9)	3244.0 (1929.7, 5803.5)	10.2 (9.4, 11.1)
Self-excitatory	Spatial lengthscale (km)	11.1 (10.1, 12.0)	23.3 (22.2, 24.4)	6.5 (5.9, 7.2)
	Temporal lengthscale (days)	1.1 (0.9, 1.4)	2.2 (1.9, 2.5)	10.0 (9.2, 10.8)
	Normalized weight	0.34 (0.31, 0.37)	0.27 (0.17, 0.36)	0.44 (0.41, 0.47)

Big model

2014-2016 Ebola virus outbreak in West Africa



▶ 1,610 sequenced viruses (1,367 of which have locations data)

► 21,811 unsequenced cases

Variable degrees of contagion

One can tailor the triggering function to change for each observation (Schoenberg et al., 2019):

$$\lambda(\mathbf{x},t) = \mu(\mathbf{x}) + \sum_{t_n < t} g_n(\mathbf{x} - \mathbf{x}_n, t - t_n).$$

In the following, I specify

$$\xi(\mathbf{x},t) = \frac{\theta_0 \omega}{h^D} \sum_{t_n < t} \theta_n \, e^{-\omega \, (t-t_n)} \phi\left(\frac{\mathbf{x} - \mathbf{x}_n}{h}\right)$$

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Phylogenetic Hawkes process



Random walk Metropolis



Hamiltonian Monte Carlo



Hamiltonian Monte Carlo

Augment parameter space with auxiliary Gaussian variable p and construct a Hamiltonian energy function:

$$egin{aligned} \mathcal{H}(\mathsf{z},\mathsf{p}) &= -\log(\pi(\mathsf{z}) imes \phi(\mathsf{p})) \ & \propto -\log\pi(\mathsf{z}) + rac{1}{2}\mathsf{p}^{\mathsf{T}}\mathsf{M}^{-1}\mathsf{p}\,. \end{aligned}$$

New states of the Markov chain are proposed by forward integrating Hamilton's equations:

$$\begin{aligned} \frac{\mathrm{d} \mathbf{z}}{\mathrm{d} t} &= \frac{\partial H}{\partial \mathbf{p}} = \mathsf{M}^{-1} \mathsf{p} \\ \frac{\mathrm{d} \mathbf{p}}{\mathrm{d} t} &= -\frac{\partial H}{\partial \mathbf{z}} = \nabla \log \pi(\mathbf{z}) \; . \end{aligned}$$

Numerical simulation induces discretization error, which we correct with a Metropolis accept-reject step.

Hamiltonian Monte Carlo

Benefits:

► HMC scales to tens-of-thousands of parameters.

Challenges:

- ► HMC necessitates repeated computation of log-likelihood and its gradient (best case O(N)).
- Preconditioning required for ill-conditioned posteriors. May involve *additional* expensive Hessian evaluations.

HMC for variable rates?

• The Hawkes likelihood scales
$$\mathcal{O}(N^2)$$
 \checkmark

• But the Hawkes log-likelihood gradient also scales $\mathcal{O}(N^2)$

$$\begin{split} \frac{\partial \ell}{\partial \theta_n} &= -\frac{\partial \Lambda_n}{\partial \theta_n} + \sum_{t_n < t_{n'}} \frac{1}{\lambda_{n'}} \frac{\partial \lambda_{n'n}}{\partial \theta_n} \\ &= \theta_0 \left(e^{-\omega \left(t_N - t_n \right)} - 1 \right) + \sum_{t_n < t_{n'}} \frac{1}{\lambda_{n'}} \frac{\theta_0 \omega}{h^D} e^{-\omega \left(t_{n'} - t_n \right)} \phi \left(\frac{\mathsf{x}_{n'} - \mathsf{x}_n}{h} \right) \end{split}$$

 And the Hawkes log-likelihood Hessian diagonal also scales *O*(*N*²)!

$$\mathsf{M}_{mm}^{-1} \approx -\frac{\partial^2 \ell}{\partial \theta_m^2} = \sum_{t_m < t_n} \frac{1}{\lambda_n^2} \frac{\theta_0^2 \omega^2}{h^{2D}} e^{-2\omega (t_n - t_m)} \phi^2 \left(\frac{\mathsf{x}_n - \mathsf{x}_m}{h}\right) \,.$$

$\left(\frac{\partial \ell}{\partial \theta_1} \right)_1$	$\left(\frac{\partial \ell}{\partial \theta_1}\right)_2$		$\left(\frac{\partial\ell}{\partial\theta_1}\right)_N$	$rac{\partial \ell}{\partial heta_1}$
$\left(\frac{\partial \ell}{\partial \theta_2} \right)_1$	$\left(\frac{\partial \ell}{\partial \theta_2}\right)_2$		$\left(\tfrac{\partial\ell}{\partial\theta_2}\right)_N$	$\frac{\partial \ell}{\partial \theta_2}$
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			•	•
$\left(\tfrac{\partial \ell}{\partial \theta_M} \right)_1$			$\left(\tfrac{\partial\ell}{\partial\theta_M}\right)_N$	$\frac{\partial \ell}{\partial \theta_M}$

$\left(\frac{\partial\ell}{\partial\theta_1}\right)_1$			
$\left(\frac{\partial \ell}{\partial \theta_2}\right)_1$			
:		·	•••
$\left(\frac{\partial \ell}{\partial \theta_M}\right)_1$			











+1				
+2				
:		·		:
+_M			•	•





Posterior inference

Generate 100 million Markov chain states (~ 3.5 million samples/day on Nvidia GV100) in 1 month.



Inferred rates of contagion



Inferred rates of contagion



Viral observation

Biologically modulated rates



There are more challenges:

- ► Flexible models (the irony of model based nonparametrics)
- Boundary issues (censoring and truncation)
- Differential sampling

Acknowledgements

Joint work with

- ► Marc Suchard (UCLA)
- ► Xiang Ji (Tulane)

Funded by K-Award *Big Data Predictive Phylogenetics with Bayesian Learning*.



Thank you!