

# MCMC with Multiple Proposals

Andrew Holbrook

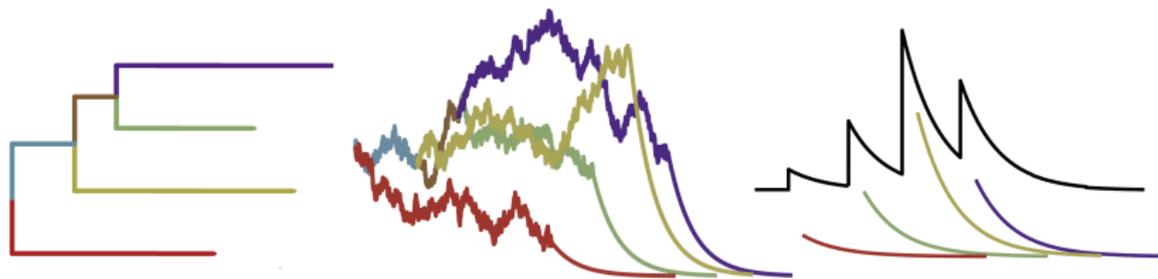
UCLA Biostatistics

May 18, 2023

# At the Intersection of Big Data and Big Model

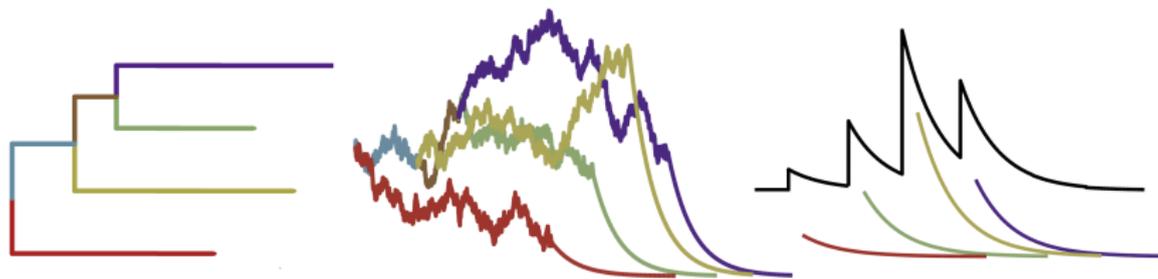
# A Unified Model for Viral Spread

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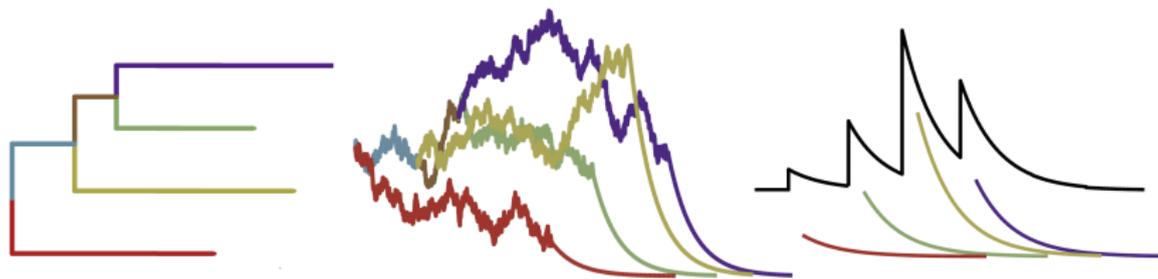
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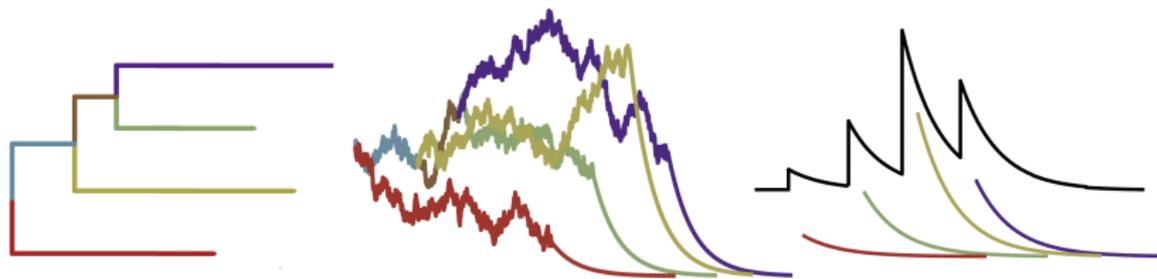
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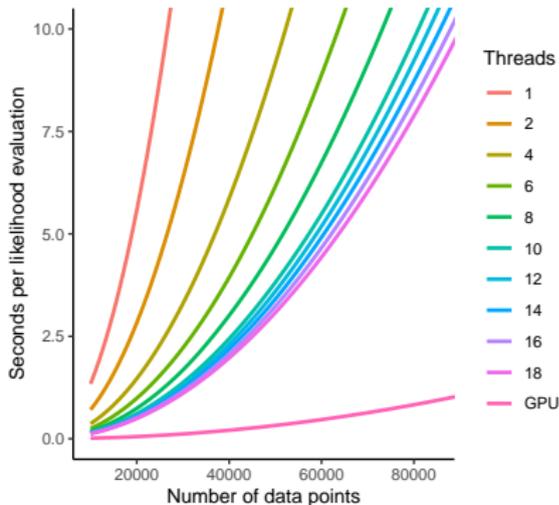
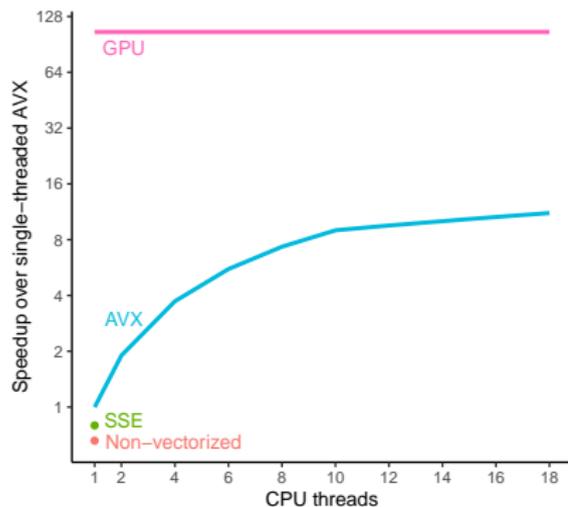
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- ▶ Virus-specific latent variables connect a spatiotemporal Hawkes process model with a phylogenetic diffusion prior.
- ▶ The number of latent variables is  $\mathcal{O}(N)$ , for  $N$  the number of observed viruses.
- ▶ Hawkes likelihood computations require  $\mathcal{O}(N^2)$  floating-point operations.



# A Unified Model of Viral Spread

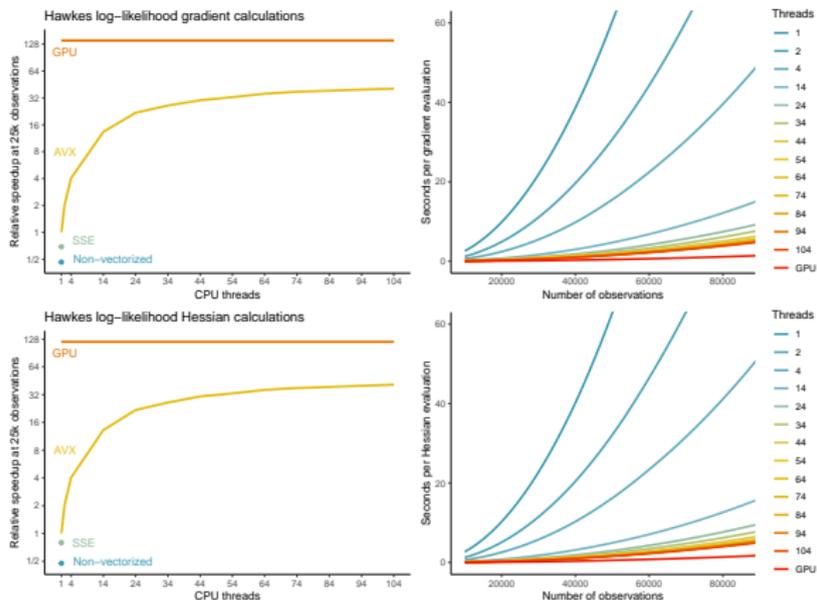
Proposal #1: use parallel computing to accelerate MH bottleneck, i.e., likelihood computations.



What about high dimensionality?

# A Unified Model of Viral Spread

Proposal #2: also use parallel computing to accelerate adaptive HMC bottlenecks, i.e., log-likelihood gradient/Hessian.

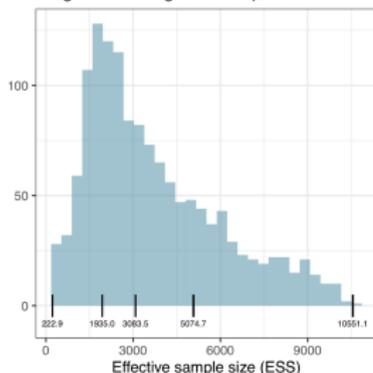


What about bad geometry (non-linearity, multimodality)?

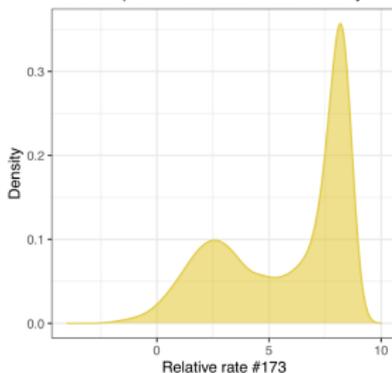
# A Unified Model of Viral Spread

Proposal #3: to analyze over 23k Ebola cases (2014-2016 West Africa), run the chain for 30 days using Nvidia GV100 GPU.

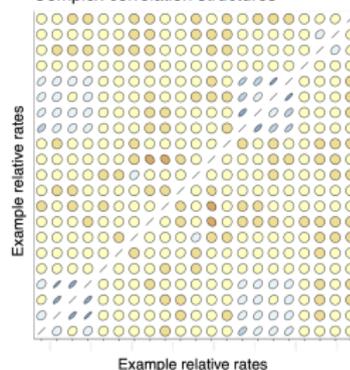
Diagnostic histogram and quartiles



An example of multiscale multimodality



Complex correlation structures



# Inexhaustive Taxonomy of Parallel MCMC

- ▶ Between-chain parallelization: multiple independent chains;
- ▶ Within-chain parallelization:
  - ▶ Model-dependent parallelization: likelihood computations;
  - ▶ Model-independent parallelization:
    - ▶ Parallel tempering;
    - ▶ Multiple-try metropolis;
    - ▶ **Multiple proposals, single acceptance step.**

## MCMC with Multiple Proposals

# A Complicated Landscape

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- ▶ Schwedes and Calderhead (2021). *Rao-Blackwellized parallel MCMC*, AISTATS. (ML), (“Parallel MCMC”)

# A Complicated Landscape

- ▶ Interesting gap between 1977 and 2003
- ▶ Literature is interdisciplinary
- ▶ Non-negligible preprint count
- ▶ A large amount of redundant, contradictory terminology
- ▶ Much of literature focuses on weighted averages (and calls this Rao-Blackwellization)
- ▶ “Parallelizable” is often conflated with “Parallelized”
- ▶ I have probably not included your work

## Efficient Multiproposal Structures

# Multiproposal MCMC

A parallel MCMC algorithm builds a transition kernel  $P(\boldsymbol{\theta}_0, d\boldsymbol{\theta})$  by:

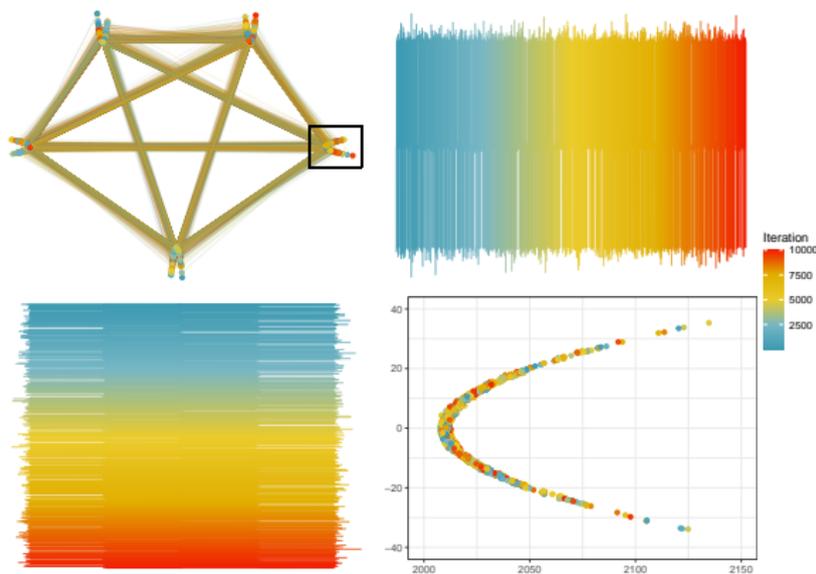
1. generating  $P$  proposals  $\boldsymbol{\Theta}_{-0} = (\boldsymbol{\theta}_1, \dots, \boldsymbol{\theta}_P)$  from a joint distribution  $Q(\boldsymbol{\theta}_0, d\boldsymbol{\Theta}_{-0}) =: q(\boldsymbol{\theta}_0, \boldsymbol{\Theta}_{-0})d\boldsymbol{\Theta}_{-0}$ ; and
2. selecting the next state with probabilities

$$\pi_p = \frac{\pi(\boldsymbol{\theta}_p)q(\boldsymbol{\theta}_p, \boldsymbol{\Theta}_{-p})}{\sum_{p'=0}^P \pi(\boldsymbol{\theta}_{p'})q(\boldsymbol{\theta}_{p'}, \boldsymbol{\Theta}_{-p'})}, \quad p \in \{0, 1, \dots, P\}.$$

This kernel maintains detailed balance and leaves  $\pi(d\boldsymbol{\theta})$  invariant.

# Multiproposal MCMC

PRO: using large numbers of proposals  $P$  helps overcome multimodality and non-linearity.



CON: requires  $\mathcal{O}(P)$  target evaluations  $\pi(\theta_p)$  and proposal evaluations  $q(\theta_p, \Theta_{-p})$ , each of the latter being  $\mathcal{O}(P)$ .

# Simplified Acceptance Probabilities

Can we somehow enforce  $q(\boldsymbol{\theta}_p, \Theta_{-p}) = q(\boldsymbol{\theta}_{p'}, \Theta_{-p'})$ ,  
 $\forall p, p' \in \{0, 1, \dots, P\}$ , to obtain simplified acceptance probabilities

$$\pi_p = \frac{\pi(\boldsymbol{\theta}_p)}{\sum_{p'=0}^P \pi(\boldsymbol{\theta}_{p'})}, \quad p \in \{0, 1, \dots, P\}.$$

Such structured multiproposals would result in  $\mathcal{O}(P^2)$  time savings and simpler implementation. I consider two such approaches in

- ▶ Holbrook (2023a). *Generating MCMC proposals by randomly rotating the regular simplex*, Journal of Multivariate Analysis.

# The Simplicial Sampler (elegant and expensive)

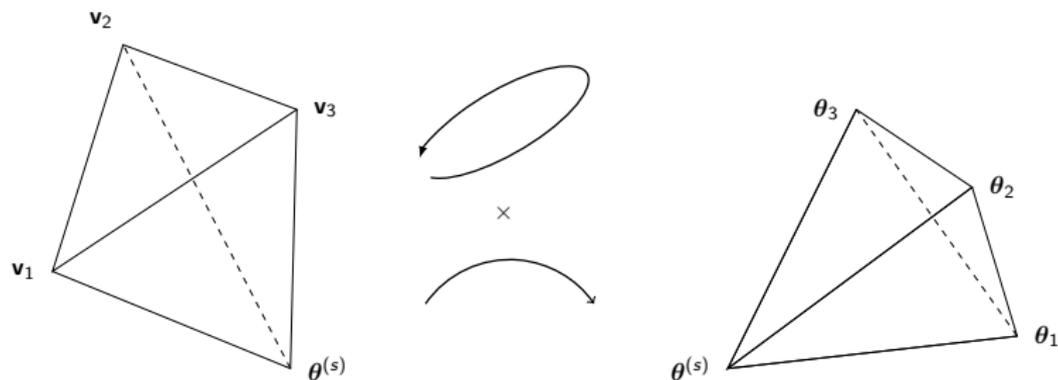
Let state space be  $\mathbb{R}^D$  and  $\mathbf{v}_1, \dots, \mathbf{v}_D \in \mathbb{R}^D$  satisfy

$$\|\mathbf{v}_d - \mathbf{v}_{d'}\|_2 = \lambda > 0, \quad d \neq d' \in \{1, \dots, D\}.$$

Then the simplicial sampler follows the following steps:

1. Sample  $D \times D$  orthonormal matrix  $\mathbf{Q}$  according to Haar distribution  $\mathbf{Q} \sim \mathcal{H}(\mathcal{O}_D)$ .
2. Rotate and translate the simplicial vertices  $(\mathbf{0}, \mathbf{v}_1, \dots, \mathbf{v}_D) \mapsto \mathbf{Q}(\mathbf{0}, \mathbf{v}_1, \dots, \mathbf{v}_D) + \boldsymbol{\theta}^{(s)} =: (\mathbf{0}, \boldsymbol{\theta}_1, \dots, \boldsymbol{\theta}_D)$ .
3. Draw a single sample  $\boldsymbol{\theta}_d$  from  $(\boldsymbol{\theta}_0, \dots, \boldsymbol{\theta}_D)$  with probability proportional to  $\pi(\boldsymbol{\theta}_d)$ .
4. Set  $\boldsymbol{\theta}^{(s+1)} = \boldsymbol{\theta}_d$ .

# The Simplicial Sampler (elegant and expensive)



A simplicial sampling multiproposal for  $D = 3$ . Proposal set is obtained by rotating three simplex vertices about current state  $\theta^{(s)}$ .

PRO: saves  $\mathcal{O}(D^2)$  time for  $D$  evaluations  $q(\theta_d, \Theta_{-d})$ .

CON:  $P = D$  and cost is  $\mathcal{O}(D^3)$ .

## Tjelmeland Correction (a free lunch)

Tjelmeland (2004) suggests the two-step multiproposal

1.  $\bar{\theta} \sim N_D(\theta^{(s)}, \Sigma)$ ;
2.  $\theta_1, \dots, \theta_P \stackrel{iid}{\sim} N_D(\bar{\theta}, \Sigma)$ .

Why? No satisfactory explanation. But it turns out that this structure leads to the desired equality (Holbrook 2023a):

$$q(\theta_p, \Theta_{-p}) = q(\theta_{p'}, \Theta_{-p'}), \forall p, p' \in \{0, 1, \dots, P\}.$$

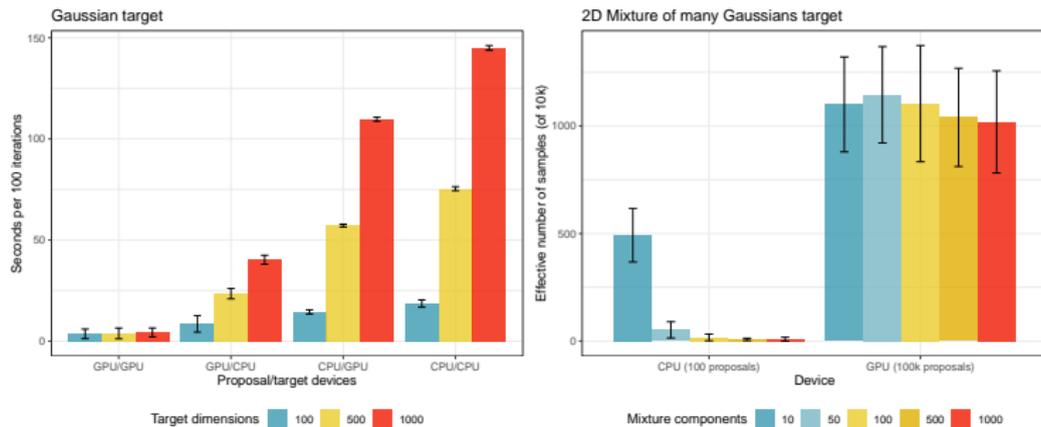
As promised, the resulting acceptance probabilities are:

$$\pi_p = \frac{\pi(\theta_p)}{\sum_{p'=0}^P \pi(\theta_{p'})}, \quad p \in \{0, 1, \dots, P\}.$$

Only the  $\mathcal{O}(P)$  target evaluations remain in our way.

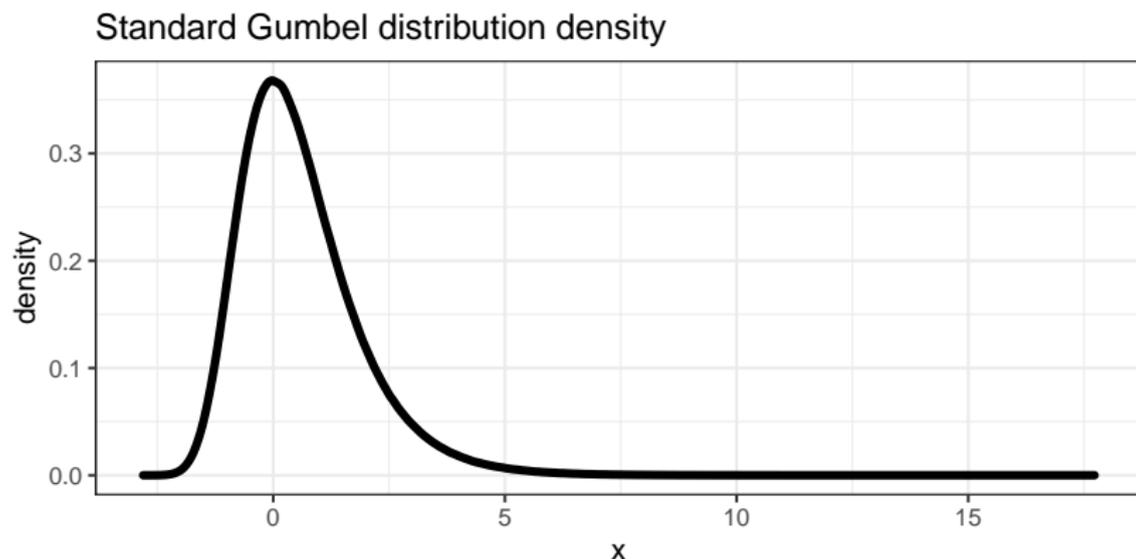
## Parallelizing Parallel MCMC

# Parallelizing Target Evaluations: CPU vs GPU



- ▶ Glatt-Holtz et al. (2022). *Parallel MCMC algorithms: theoretical foundations, algorithm design, case studies*, Preprint.

# The Gumbel Distribution



If  $z \sim \text{Gumbel}(0, 1)$ , then it has density and distribution functions

$$g(z) = \exp(-z - \exp(-z)) \quad \text{and} \quad G(z) = \exp(-\exp(-z)).$$

# Gumbel-Max Trick

We wish to sample from the discrete distribution  $\hat{p} \sim \text{Discrete}(\pi)$  for  $\hat{p} \in \{0, 1, \dots, P\}$  and we only know  $\pi^* = c\pi$  for some  $c > 0$ .

Define  $\lambda^* = \log \pi^* = \log \pi + \log c$  and suppose  $z_0, z_1, \dots, z_P \stackrel{iid}{\sim} \text{Gumbel}(0, 1)$ .

Finally, define  $\alpha_p^* := \lambda_p^* + z_p$  and  $\hat{p} = \arg \max_{p=0, \dots, P} \alpha_p^*$ .

Then the following holds (Papandreou and Yuille, 2011):

$$\Pr(\hat{p} = p) = \pi_p, \quad p = 0, 1, \dots, P.$$

**Data:** Initial Markov chain state  $\theta^{(0)}$ ; total length of Markov chain  $S$ ; total number of proposals per iteration  $P$ .

**Result:** A Markov chain  $\theta^{(1)}, \dots, \theta^{(S)}$ .

**for**  $s \in \{1, \dots, S\}$  **do**

$\theta_0 \leftarrow \theta^{(s-1)}$ ;

$\bar{\theta} \leftarrow \text{Normal}_D(\theta_0, \Sigma)$ ;

$z_0 \leftarrow \text{Gumbel}(0, 1)$ ;

**for**  $p \in \{1, \dots, P\}$  **do**

$\theta_p \leftarrow \text{Normal}_D(\bar{\theta}, \Sigma)$ ;

$z_p \leftarrow \text{Gumbel}(0, 1)$ ;

**end**

$\hat{p} \leftarrow \arg \min_{p=0, \dots, P} \left( f(p) := -(z_p + \log \pi(\theta_p)) \right)$ ;

$\theta^{(s)} \leftarrow \theta_{\hat{p}}$ ;

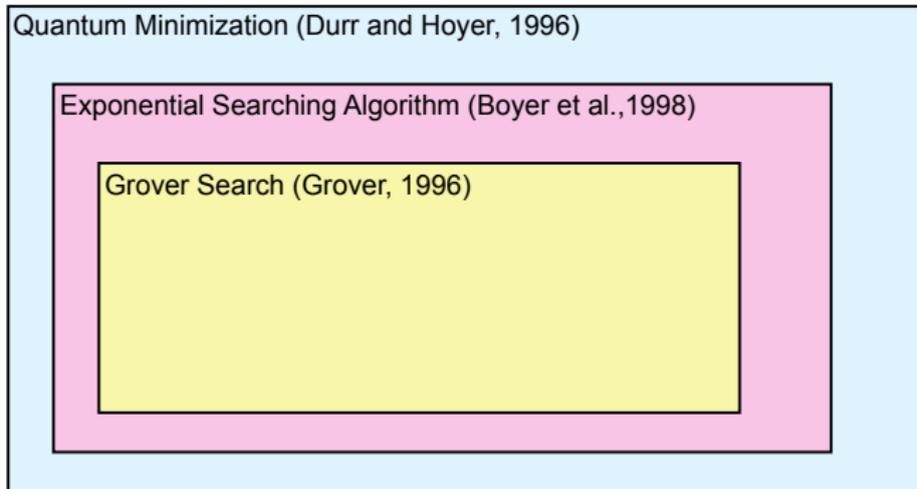
**end**

**return**  $\theta^{(1)}, \dots, \theta^{(S)}$  .

# Quantum Parallel MCMC

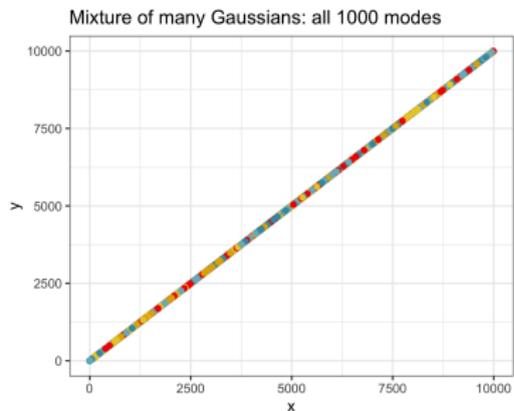
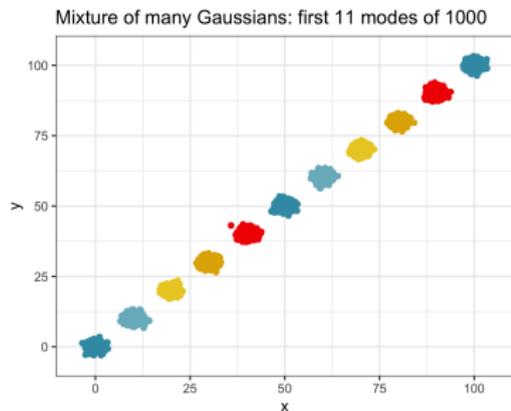
Use a quantum circuit to obtain

$$\hat{p} = \arg \min_{p=0, \dots, P} \left( f(p) := -(z_p + \log \pi(\theta_p)) \right)$$



- ▶ Holbrook (2023b). *A quantum parallel Markov chain Monte Carlo*, JCGS.

# QPMCMC: Racing to an ESS of 100



Proposals	MCMC iterations	Target evaluations	Speedup	Efficiency gain
1,000	249,398 (200,998, 311,998)	24,988,963 (20,149,132, 31,265,011)	9.98 (9.98, 9.98)	1
5,000	14,398 (12,998, 16,998)	3,314,560 (2,989,418, 3,916,281)	21.72 (21.70, 21.74)	7.58 (6.25, 9.71)
10,000	5,998 (4,998, 6,998)	1,993,484 (1,662,592, 2,330,842)	30 (29.96, 30.26)	12.87 (8.64, 18.80)

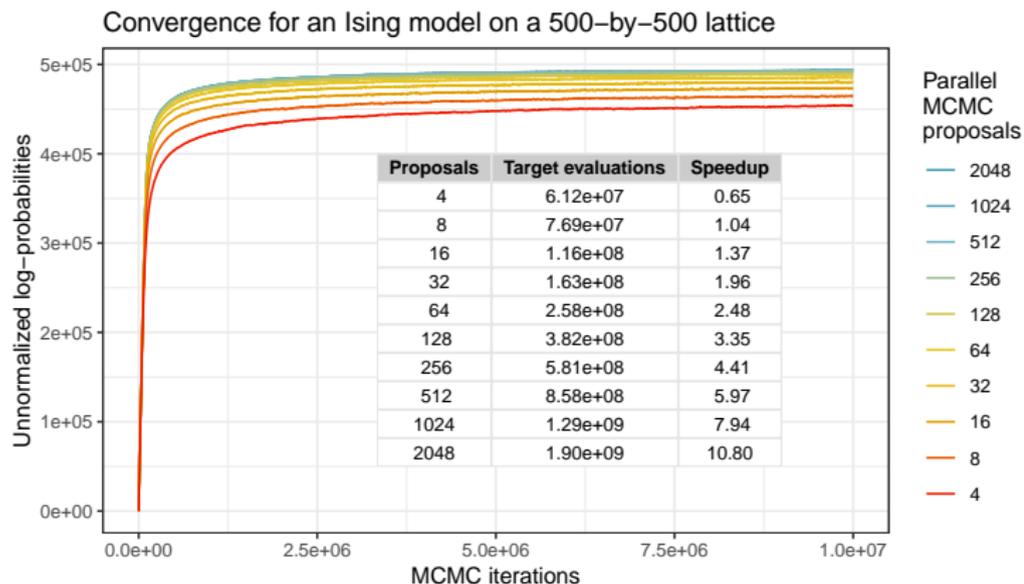
# Ising Model Target

Consider the Ising-type lattice model over configurations  $\boldsymbol{\theta} = (\theta_1, \dots, \theta_D)$  consisting of  $D$  individual spins  $\theta_d \in \{-1, 1\}$

$$\pi(\boldsymbol{\theta}|\rho) \propto \exp \left( \rho \sum_{(d,d') \in \mathcal{E}} \theta_d \theta_{d'} \right).$$

No need for Tjelmeland corrections when we use uniform proposals on  $\{-1, 1\}^D$ . The following results are based on single-flip proposals (although not necessary).

# Ising Model Target



# Bayesian Image Segmentation

Following Hurn (1997),  $y_d$  are intensity values associated with individual pixels.

$$y_d | (\mu_\ell, \sigma^2, \theta_d) \stackrel{iid}{\sim} \text{Normal}(\mu_\ell, \sigma^2), \quad y_d \in [0, 255],$$

$$\theta_d = \ell, \quad d \in \{1, \dots, D\},$$

$$\mu_\ell \stackrel{iid}{\sim} \text{Uniform}(0, 255), \quad \ell \in \{-1, 1\},$$

$$\frac{1}{\sigma^2} \sim \text{Gamma}\left(\frac{1}{2}, \frac{1}{2}\right)$$

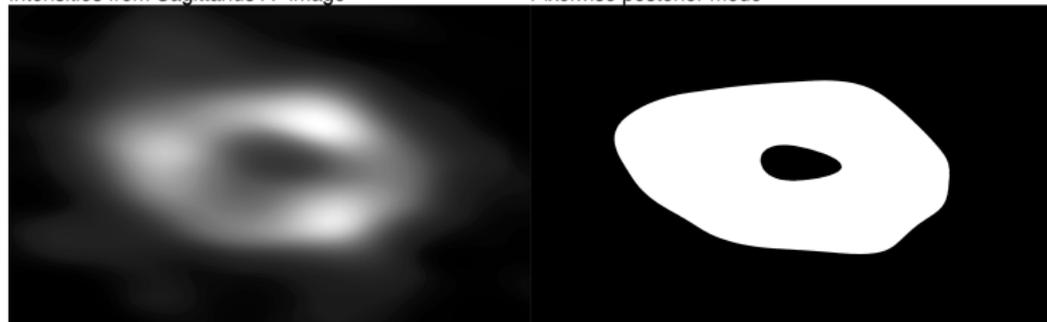
$$\boldsymbol{\theta} \sim \text{Ising}(\rho), \quad \rho = 1.2.$$

# Bayesian Image Segmentation

Segmenting a 4,076-by-4,076 intensity map. Using 1,024 proposals, QPMCMC requires less than 10% the evaluations required by a conventional computer.

Intensities from Sagittarius A\* image

Pixelwise posterior mode



## Future Directions

# Theoretical Challenges

Glatt-Holtz et al. (2022) develop foundations for multiproposal MCMC, incorporating:

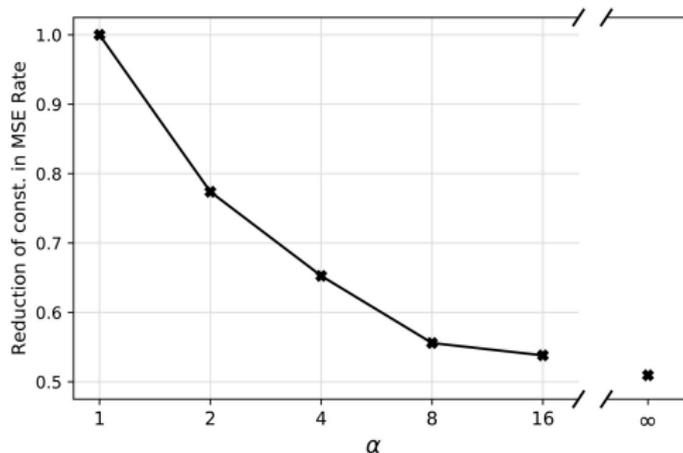
- ▶ general state space representation (Tierney, 1998);
- ▶ involutions on extended phase spaces (Nekludov et al., 2020; Glatt-Holtz et al., 2020; Andreiu et al., 2020);
- ▶ proposal cloud resampling;
- ▶ Metropolis-Hastings and Barker/Boltzmann acceptances.

We still lack:

- ▶ Optimal tuning guidances  $(D, P)$ ;
- ▶ Error bounds for biased kernels;
- ▶ nonreversible multiproposal MCMC.

# Bizarre Benefits of Bias

Schwedes and Calderhead (2021) estimate the relative reduction in MSE for Monte Carlo estimators as a function of  $\alpha$ , where  $\alpha \times P$  is the number of proposal cloud resampling iterations.



This is using the naive multiproposal

$$\theta_1, \dots, \theta_P \stackrel{iid}{\sim} N_D(0, \Sigma).$$

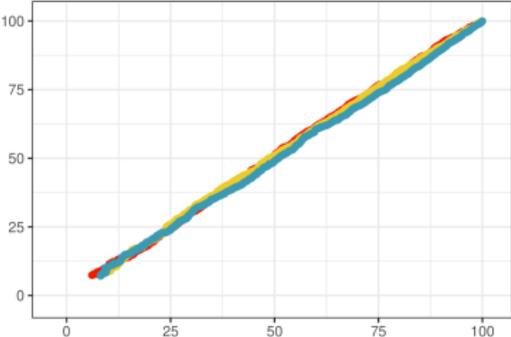
with full acceptance probabilities  $\pi_p \propto \pi(\theta_p) \prod_{p' \neq p} q(\theta_p, \theta_{p'})$ .

# Bizarre Benefits of Bias

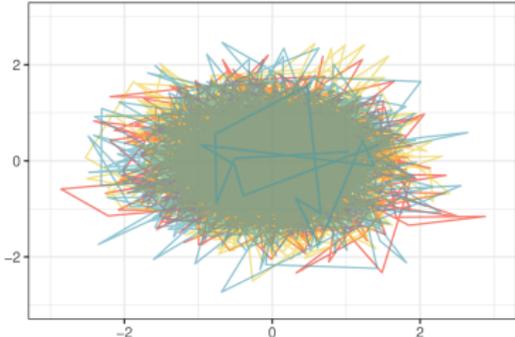
Starting correct algorithm at origin



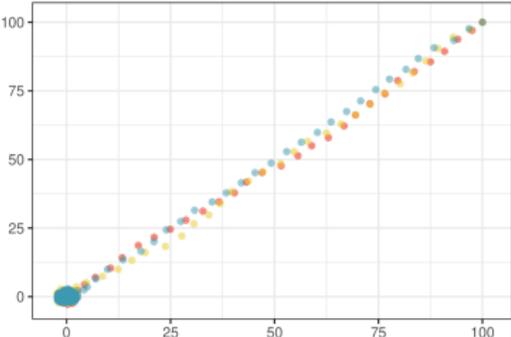
Starting correct algorithm away from origin



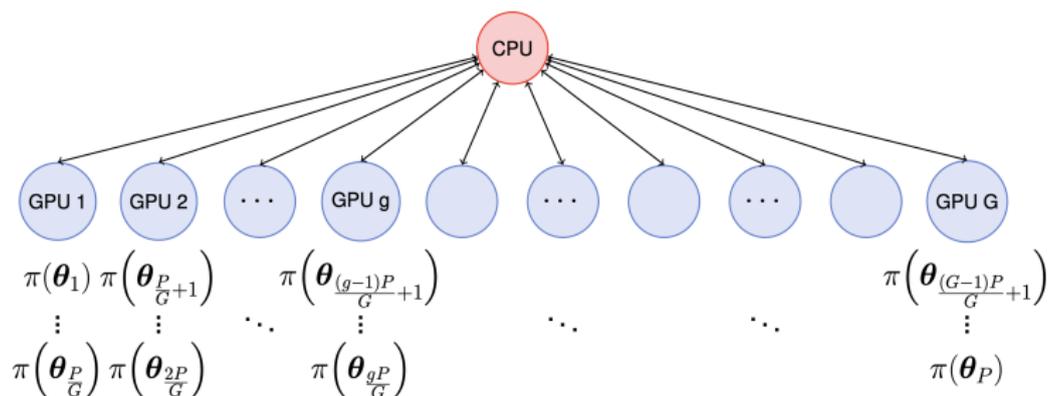
Starting an incorrect algorithm at origin



Starting an incorrect algorithm away from origin



# Tjelmeland Correction Reduces Communication



If models are multimodal and parallelizable:

- ▶ Bayesian inversion of nonlinear PDEs;
- ▶ Hawkes processes (temporal, spatiotemporal, multivariate).

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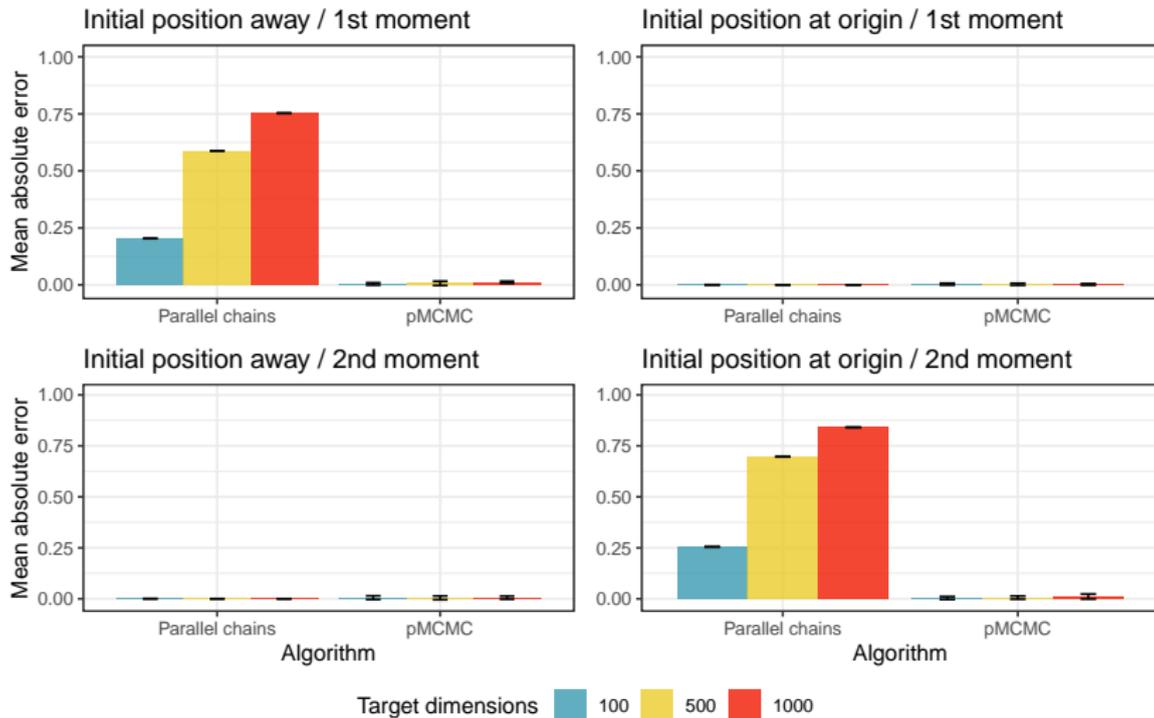
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## Many Proposals vs Many Chains

# Multivariate Gaussian Targets



# Massively Multimodal Targets

